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Section 5: Continuous Distributions

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1 Summary

1.1 Continuous Random Variables

Definition 1 (Continuous r.v.). A continuous random variable has an interval for its support.

More precisely: a continuous random variable has uncountable support, while discrete random variable shave finite/countably infinite support.

Definition 2 (Cumulative distribution function). The **cumulative distribution function (CDF)** of a random variable *X* is the function $F : \mathbb{R} : [0, 1]$ defined by $F(x) = P(X \le x)$.

★ 3. For a continuous random variable *X*, *any* value of $x \in \mathbb{R}$ (including those in the support) has P(X = x) = 0. This also means that the CDF can be defined in multiple ways since $P(X \le x) = P(X < x) + P(X = x) = P(X < x)$.

Definition 4 (Probability density function). For a continuous random variable *X* with CDF *F*, the **probability density function (PDF)** is $f : \mathbb{R} \to \mathbb{R}^{\geq 0}$ defined as

$$f(x) = \frac{d}{dx}F(x).$$
(1)

Probability density functions satisfy the following condition:

$$\int_{-\infty}^{\infty} f(x)dx = 1.$$

$$f_X(x) \neq f_{g(X)}(g(x)).$$

1.1.1 Uses of CDFs and PDFs

For *any* random variable *X* (continuous or discrete), you can use the CDF to calculate the following:

$$P(X > x) = 1 - P(X \le x) = 1 - F(x)$$

$$P(x_1 < X \le x_2) = P(X \le x_2) - P(X \le x_1) = F(x_2) - F(x_1)$$

You can assume that the CDF of a continuous random variable is differentiable, so that the PDF can actually be defined. We can find the probabilities of intervals by integrating the PDF and adjusting

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the bounds:

$$P(X \le x) = P(X < x) = \int_{-\infty}^{x} f(x)dx$$
$$P(X \ge x) = P(X > x) = \int_{x}^{\infty} f(x)dx$$
$$P(x_1 < X < x_2) = \int_{x_1}^{x_2} f(x)dx.$$

1.1.2 Continuous analogs of all of our tools

The general rules are:

- Integrals instead of sums.
- PDFs instead of PMFs.
- When asked to find the distribution of a continuous random variable, it's much easier to work with the CDF than the PDF.

So here's a table with the tools we've talked about:

Tool	Discrete	Continuous
Expectation	$E(X) = \sum_{x} x P(X = x)$	$E(X) = \int_{-\infty}^{\infty} x f_X(x) dx$
LOTUS	$E(g(X)) = \sum_{x} g(x)P(X = x)$	$E(g(X)) = \int_{-\infty}^{\infty} g(x) f_X(x) dx$
Bayes' rule	$P(X = x Y = y) = \frac{P(Y = y X = x)P(X = x)}{P(A)}$	$f_{X Y=y}(x) = rac{f_{Y X=x}(y)f_X(x)}{f_Y(y)}$

1.2 Uniform

Definition 6 (Uniform distribution). For any interval (a, b), a random variable $U \sim \text{Unif}(a, b)$ has a uniform distribution (i.e., constant PDF) over the support (a, b). There is no uniform whose support is the full real line.

The PDF and CDF can be derived:

$$f_U(x) = \begin{cases} \frac{1}{b-a} & x \in (a,b) \\ 0 & x \notin (a,b) \end{cases}$$
$$F_U(x) = P(U < x) = \begin{cases} 0 & x \le a \\ \frac{x-a}{b-a} & x \in (a,b) \\ 1 & x \ge b \end{cases}$$

2 Practice Problems

1. Suppose *X* is a random variable with the following PDF:

$$f_X(x) = \begin{cases} \frac{1}{x^2} & x \ge 1, \\ 0 & x < 1. \end{cases}$$

(a) What is the expected value of *X*?

Solution

$$E(X) = \int_{1}^{\infty} x \frac{1}{x^{2}} dx = \int_{1}^{\infty} \frac{1}{x} = \ln x |_{1}^{\infty} = \infty.$$

(b) What is the expected value of 1/X?

Solution Using LOTUS,

$$E(1/X) = \int_1^\infty \frac{1}{x} \frac{1}{x^2} dx = -\frac{1}{x^2} \Big|_1^\infty = 1.$$

(c) What is the distribution of X^2 ?

Solution

The support of X^2 is $[1, \infty)$. We can find the CDF, assuming $y \ge 1$:

$$P(X^{2} \le y) = P(X \le \sqrt{y}) = \int_{1}^{\sqrt{y}} \frac{1}{x^{2}} dx = -\frac{1}{x} \Big|_{1}^{\sqrt{y}} = 1 - \frac{1}{\sqrt{y}}$$

Given that $P(X^2 \le y) = 0$ if y < 1, we have defined our CDF. That is sufficient for defining the distribution.

If you're curious about the PDF:

$$f_{X^2}(y) = \begin{cases} \frac{1}{2y^{3/2}} & y \ge 1\\ 0 & y < 1 \end{cases}$$

(d) Now suppose *X* is the side length of a square. What is the distribution of the area of the square?

Solution

The square has area X^2 , so the distribution of the area is the same as the distribution of X^2 , as found in part (a)

2. Suppose *Y* is a random variable with the following PDF:

$$f_{\mathrm{Y}}(y) = \frac{1}{2}e^{-|y|}.$$

(a) What is the distribution of |Y|?

Solution We can derive the CDF: for $y \ge 0$ we get

$$P(|Y| < y) = P(-y < Y < y) = \int_{-y}^{y} \frac{1}{2} e^{-|x|} dx$$
$$= 2 \int_{0}^{y} \frac{1}{2} e^{-|x|} dx = \int_{0}^{y} e^{-x} dx$$
$$= -e^{-x} |_{0}^{y} = 1 - e^{-y}.$$

So the full CDF is

$$P(|Y| < y) = \begin{cases} 1 - e^{-y} & y \ge 0, \\ 0 & y < 0. \end{cases}$$

As we'll see soon, $|Y| \sim \text{Expo}(1)$.

(b) What is the expected value of *Y*, given that the expectation does exist (i.e., is not infinite)? (*hint: use symmetry*)

Solution

$$\begin{split} E(Y) &= \int_{-\infty}^{\infty} \frac{1}{2} y e^{-|y|} dy \\ &= \int_{0}^{\infty} \frac{1}{2} y e^{-y} dy + \int_{-\infty}^{0} \frac{1}{2} y e^{y} dy \\ &= \frac{1}{2} \int_{0}^{\infty} y e^{-y} dy + \frac{1}{2} \int_{0}^{\infty} -y e^{-y} dy \end{split}$$

The two terms are identical up to a negative, so they actually cancel out to give E(Y) = 0.

(c) What is the expected value of $e^{-|Y|}$? Use LOTUS on the distribution of |Y| that you found in part (a).

Solution

If we let Z = |Y|, then we find that the PDF is $f(z) = e^{-z}$ for $z \ge 0$ and f(z) = 0 for z < 0. We can then use LOTUS:

$$E(e^{-Z}) = \int_0^\infty e^{-z} e^{-z} dz = \int_0^\infty e^{-2z} dz$$

We know that $\int_0^\infty e^{-z} dz = 1$ since Z has a valid PDF. So let's set u = 2z and u-

substitute:

$$E(e^{-Z}) = \frac{1}{2} \int_0^\infty e^{-u} du = \frac{1}{2}$$

- 3. Suppose $X \sim \text{Unif}(0, 1)$.
 - (a) Find P(X = 0.3).

Solution

Since *X* is continuous, P(X = 0.3) = 0.

(b) Find P(X < 0.3).

Solution

Using the CDF, P(X < 0.3) = 0.3.

(c) Find P(0.3 < X < 0.7).

Solution

We can use the CDF:

$$P(0.3 < X < 0.7) = P(X < 0.7) - P(X < 0.3) = 0.7 - 0.3 = 0.4.$$

We could also integrate the PDF:

$$P(0.3 < X < 0.7) = \int_{0.3}^{0.7} 1 dx = 0.7 - 0.3 = 0.4.$$

(d) Find $P(0.3 \le X \le 0.7)$.

Solution

This is the same as the previous part since the extra equalities don't matter.